

Abstract

This study examines whether household access to microfinance reduces poverty, and if so, to what extent and acros s which dim ensions of wellbeing. The study draws on first-hand observations and empirical dat a gathered from interviews of 1,132 households across 11 districts in the rural areas of the province of Punjab in Pakistan. It employs a quasi-exp erimental research design and makes use of data collected b y int erviewing bot h borrower (t reatment) and non-borrower (control) ho useholds. Sample select ion biases are cont rolled b y ma tching propensity s cores. Findings reveal that al though borrowers seem to fare better than non-borrowers a cross around 70 percent of the indicat ors, a majority of these are n ot statistically significant. This suggests that despite producing some degree of posit ive impact, microfinance institutions still have t o make s ustained efforts to bring about real difference to the livelihoods of the poor.

Keywords: microfinance, poverty, impac t assessment, propensity score matching, Pakistan

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1. Introduction

Poor households in urban and, in particular, rural areas in many developing countries do not have easy access t o basic f inancial services. Their 'syst ematic exclusion ' f rom f ormal financial se rvices ha s led t o t he evolut ion of an alt ernative mod e of finance called microfinance, where financial services are provided not through traditional routes, such as local money lenders, cooperatives or banks, but through NGOs or microf inance institutions (MFIs). Microfinance has evolved and expan ded f rom Bangladesh t o ot her de veloping countries over t he last three decad es, based o n the conviction that the livelihoods of such financially-excluded poor households, without any physical collateral or credit history, can be improved if t hey have access t o small-scale loans or ot her financial services, such as savings or insurance, offered either to a group or to individuals.

The concept and practice of microfinance, however, has changed dramatically over the last decade as the microfinance sector increasingly adopts a financial systems approach, either by operating on commercial line s or by systematically reducing reliance on interest rate subsidies and/or aid ag ency financial support (Hulme and Arun 2009). As oppose d to the 'welfarist' or poverty a pproach, the 'self -sustainability' or 'f inancial sy stems' approach advocated by the institutionists now covers mainly non-poor or relatively less-poor clients on the fringes of the formal financial system, and it does not target the poorest because of the need for the financial sustainability of the MFIs themselves. As MF Is are required by this approach to lessen t heir reliance on donor f unds and subsidies and adopt good banking practices, they are expected to innovate to ensure provision of more efficient and better financial services with lower costs. Profits are vie wed as being not only acceptable, but essential, because t hey are expect ed to a ttract private investment to the sector (Conning 1999). Whilst many MFIs have begun t o place more emphasis on the financial systems approach d uring t he recent global recession, some of the major MF Is have designed specialised and targeted products for the very poor. For example, Grameen Bank and BRAC in Bangladesh offer financial products specifically tailored and targeted at the needs of the ncome Generat ion for Vulnerable Groups Developme poorest. BRACs I nt (IGVGD) programme, 'provides f ood subsidies and int ensive skills t raining to vulnerable women, as well as a standard package of microcredit, healthcare and social services' (Maes and Foose 2006, p.11).

While a few empirical studies at the micro level have shown that participants in microfinance programmes have progressive ly become capable of a ccessing f inancial ser vices and escaping from poverty (Matin et al. 2008, Hossain and Zahra 2008), the wider literature on impact evaluations of large-scale pr ogrammes has revealed mixed and conflicting findings, with some d isagreements amongst academics and practitioners about the effectiveness of microfinance as a poverty reduction measure. At one end of the spectrum lie the studies that have concluded that microfinance is a posit ive and effective measure of poverty reduction (e.g. Hossain 1988; Barnes 2001; Dunn 2002; Snodgrass and Sebst ad 2002; Goldberg 2005; Khandker 2005; Rabbani et al. 2006; Haseen 2006; Mahjabeen 2008; Banerjee, Duflo et al. 2009; Imai et al. 2010). At the opposit e end are st udies which have argu ed t hat employing this strategy has in fact driven people into greater poverty and has weakened the position of women even further, rather than empowering them (e.g. Goetz and Gupta 1996;

Neff 1996; George 2006; Chanana 2007; Bateman 2008). In between, there are studies that have caut ioned against consider ing microf inance as a 'cure-all', yet have endor sed it as assisting people to a cert ain extent, and have urged t hat it should be used with 'cautious optimism' (e.g. Bello 20 06; Banerjee et al. 2009; Karlan and Zinman 2009). Regardless of the different and apparent ly contradictory conclusions that have been derived f rom these empirical studies, which might reflect the diverse settings of the studies (focusing on different geographical areas or drawing on different methodologies), impact assessment nevertheless remains one of the most powerf ult ools by which programme e ffectiveness can be measured.

In Pakistan, the microfinance sector has been operational in various forms and sizes for over four decade s. Nevert heless, there is a deart h of reliable s tudies a ttempting to measure impact using rigorous methods. C laims about t he impact of microfinance are not well documented or support ed by verifiable evidence (Hussein and Husse in 2003), one of the main reasons for this being the limited availability of primary and secondary data in Pakistan (OPM, 2006).

There are, however, a few empirical studies that have generally confirmed that microfinance intervention has had so me positive impacts on the welfare of households in Pakistan. For example, Hussain (2003) shows that there are significant differences between participants and non-part icipants in microf inance programmes in t erms of mon thly per capit a expenditure, living cond itions, literacy rates and, more importantly, in crease in in come of participants. Mont gomery (2005) cont ends t hat microcredit programmes have a positive impact on b oth economic and social indicat ors of welf are, as well as income-generat ing activities, especially for the very poorest participants in the programme. Finally, Shirazi and Khan (2009) show t hat micro finance programmes have a posit ive impact on poverty reduction in Pakistan and argue that borrowers tend to shift to higher income groups durin g the given period. In contrast to Montgomery's findings, however, they show that the poverty status of extremely po or borrowers increase s only marginally, as they believe t hat the chronically poor borrow essentially for protection, as opposed to investing in entrepreneurial activities. There is no conclusive evidence of the impact of microfinance in Pakistan and the present st udy is one of the few to evaluate microf inance programmes where sample selection bias is controlled for.

Multi-dimensional a spects of poverty are particularly relevant t o Pa kistan. The poor in Pakistan not only have low levels of income, they also lack access to basic services, such as clean drinking water, adequate sanitation, proper education, financial services, employment opportunities, efficient markets, and sufficient and timely health facilities (World Bank, 2007). Despite con siderable ef forts t hrough various p overty alleviat ion programmes, widespread social and e conomic poverty remains a core pr oblem in Pakistan, as its economy is based predominantly on agricult ure. Almost 65 percent of the population reside in rural are as and are directly or indirectly linked to agriculture (CIA 2010, World Bank 2002). The FAO (2009) estimates that around 6 6 percent of the population of Pakistan relies on agricult ure for its livelihood. Consequent ly, the poor are over whelmingly concentrated in rural area s, where the poverty headcount is 27 percent, more than double t he size o f t hat in urb an areas. Furthermore, 80 percent of the total poor live in rural area s (IMF 2010). According to the

2007-08 estimates, 22.3 percent of the country's population lives below the poverty line, with another 20.5 percent living in vulnerable conditions (Haq 2008).

As there are no off icially-published poverty figures for Pakistan for 2009, researchers have estimated these at various levels. Ahmed an d Donoghu e (2010), for inst ance, est imate poverty to have climbed to as much as 40 perc ent, an increase of almost 80 perc ent from the 22 percent recorded in 2006. Given t he poor perf ormance that the country showed in terms of GDP growt h rate (onl y 1. 2 percent in 2009), coupled with t he high inflation experienced during 2008-09 (22 percent) and t he country's involvement in internal and external conflicts, est imates such as t hese cannot be regarded as excessive. The recent flooding in the country will place an additional burden on the already fragile economy and, as analysts say, will drag the country back by many years. Given these signs, poverty levels are set to rise in the coming years, and the targets and growth forecasts seem over-ambitious.

The limited access to financial services in the developing world is one of the main obstacles to both income generation and social prot ection. Nenova et al. (2009: ix) report that out of the 40 percent of the Pakistani population which does not engage in eit her the formal or informal financial system, an est imated 19 percent have volunt arily excluded t hemselves through lack of underst anding, a wareness, need, or for religiou s reasons. Despit e considerable efforts, microfinance has been slow to scale up, and out reach to women has been especially limited. It is est imated that o nly about eight percent of poor ho useholds receive credit from formal sources (World Bank 2007). The size of Pakistan's population and the nu mber of the poor imply that there is a large pot ential market f or microfinance in Pakistan, which, according to PMN estimates, is close t o 27 million ind ividuals (Haq 2008), thus bringing the current penetration rate to just 6.97 percent.

The rest of this paper is organised as f ollows. The next section summarises the survey design and descriptive statistics. Section 3 describes the econometric methodology and model used to control for sample selection biases. Section 4 discusses the results obtained and the main findings of the study. The concluding remarks are presented in Section 5.

2. Survey design and data

This study aims to assess the nature, extent and direction of the socio-economic impact of microfinance programmes on b orrowers, based on det ailed cross-sect ional prima ry household surveys conducted over 11 dist ricts across the rural parts of Punjab, in Eastern Pakistan. The study is based on a quasi-experimental design survey,¹ whereby comparison is made be tween t wo groups of respondent s: t he cont rol group (rep resented b y non-borrowers) and t he treatment gro up (comprising borrowe rs). The total sample of 1,132 respondents comprises 463 borrowers and 669 non-borrowers. The hypothesis that we test in our st udy is: *participation in microfinance programmes im proves the socio- economic conditions of member households*.

¹ The field survey was carried out by one of the authors between 2008 and 2009. The questionnaire and further details of the survey will be furnished on request.

In order to select households, a f our-stage random st ratified sampling t echnique was applied. In the first stage, 11 out of the 36 districts were selected from the entire province. Districts were selected systematically, as oppo sed to being select ed randomly, in order t o control for social and economic disp arities that occur across the province between various districts, and to ensure that the selected districts represent maximum and diverse population across the entire province. Starting from the north of the province, districts were select ed towards the east, west and south. In the second st age, at least one *tehsil*² was ran domly selected f rom each ident ified di strict. In the t hird st age, a t least two villag es wer e subsequently selected randomly from amongst the selected tehsils, and in t he fourth and final st age participating and non-part icipating households were sele cted at random for conducting surveys.

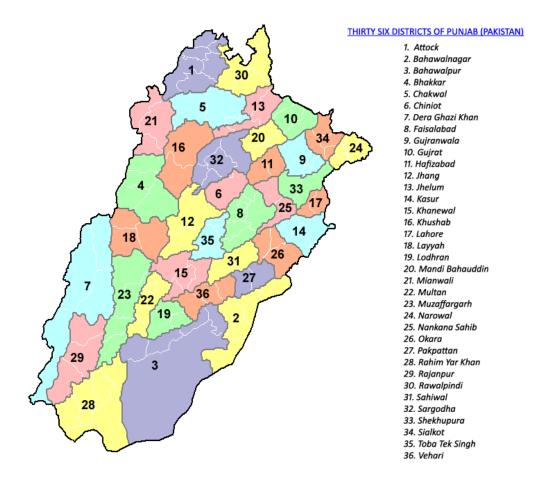


Figure 1: Map of Punjab showing the 36 districts of the province

² For admini strative purposes, Paki stan is divided in to four provin ces and a F ederal Capital. Each province comprises several districts, further divided into *tehsils* as administrative divisions. As entities of the local government, tehsils exercise certain fiscal and administrative powers over the villages and municipalities within their jurisdiction.

2.1 Selection and choice of indicators applied

Due to the multidimensional nature of poverty (Armendariz and Morduch 2005; Daley-Harris 2006), it is necessary to have a representative nature of dimensions and accompan ying indicators that would ref lect the a ctual poverty of a typical household wit hin the sample frame. After careful screening and extensive pilot testing, the final field instrument comprised questions designed t o capt ure information across t he following f our dimensions: human resources; dwell ing; f ood securit y and vulnera bility; and ownership of household asset s. Table 1 lists the dimensions and related indicators used in the survey.

Human resources	Dwelling-related indicators	Food security and vulnerability	Ownership of household assets
Age and sex of adults in household Adult literacy Number of children Occupations of adults in household Number of children below the age of 15 in household Annual expenditure on clothing and footwear for all members in household	 House ownership Type of floor Material used for constructing exterior walls and roof Number of rooms in the house Source of water supply Type of toilet. Method of bathroom waste disposal Energy for I ighting in the house Type of fue I used for cooking Structural condition of house 	Number of days when staple foods were served Number of days when vegetables were served Number of days when meat was served	Livestock (cattle and buffalo, sheep and goats, poultry, horses and donkeys, etc.) Transportation- related assets (motorcycle, bicycle, carts) Appliances and electronics (television, VCR, refrigerator, washing machine, radio/tape/stereo, mobile phone, sewing machine, etc.)

Table 1: Final list of dimensions and related indicators used in survey

The questionnaire was field-tested and a number of indicators were consequently altered to control for local specificities, and to ensure that they fully captured and reflected the relative poverty levels of both groups of households. Indicators such as t hose relating to highly contextual and subject ive responses were subsequent ly dropped from the final field instrument.

2.2 Descriptive statistics and explanation of variables

The survey respondents represented eight MFIs in the province. As shown in Appendix 1, given the strong nationwide presence of the National Rural Support Programme (NRSP), its borrowers represented almost 32 p ercent of the total sample. Kash f Foundation's strong presence and extensive outreach in the districts surrounding the provincial capital gave it a share of 28 percent, and the Punjab Rural Support Programme (PRSP) was represented by 14 percent of t hose int erviewed. Appendix 2 ref lects the number of loan cy cles t hat respondents had complet ed at the time of in terview. Almost 60 percent were f ound to be within their f irst two years of borrowing, while 16 percent were in their third cy cle. By principal occupation, alt hough the largest group of respondent s were involved in casual labour, at over 32 percent, there is a significant disparity when data is disaggregated across borrowers and non-borrowers. That is, 22 percent of borrowing hou seholds report ed their occupation as casual labour, as opposed to almost 40 percent of non-borrowing households.

For social and cultural reasons, extended families are common in Paki stan, particularly in the rural areas. The most commonly occurring size of households (mode) was five members. The mean size calculat ed from the data was 5.98 members per household and the median value 6.00. Household sizes of five to seven members constituted almost 50 percent of the entire sample, while t hose consisting of eight or more me mbers amounted to around one quarter, and single t o four-member household size is 6.58 members, according t o the Household I ntegrated Economic Survey (GoP 2009a), while t he average f or Pun jab was reported as 6.33 members for 2007-08, close to the mean (5.98) and median (6.00) values reported in the survey results. Household sizes are illustrated in Appendix 3.

In terms of loan size, as shown in Appendices 4 and 5, 22 percent of respondents had availed themselves of loans ranging from Rs.5,000 to Rs.10,000, and 30 percent had credit facilities ranging f rom Rs. 11,000 t o Rs. 15,000. Taken t ogether, these loans (up to Rs.15,000) constituted more than half of the sample. Instalment amounts also corresponded proportionately to the size of loans; it was noted t hat over 60 percent of the instalment amounts varied from Rs.1,000 to Rs.2,000, followed by smaller amounts of up to Rs.1,000, and larger amounts that ranged from Rs.2,000 to Rs.2,500, accounting for almost a quarter of the total sample. The sample mean is Rs.17,473, and the median value Rs.15,000.

Literacy rate, according to the Pakistan Social and Living St andards Measurement Survey (PSLM) for 2007-08 (for both males and females, aged 10 and above) was 56 percent at the national level and 53 percent for rural Punjab (GoP 2009b, p. 43). Data from this survey found the adult literacy rate (household members aged 15 and above) t o be 39.92 percent, whereas it was 40. 02 percent according t o PSLM (2007 -08). UNESCO's Asia- Pacific Literacy Data Base (20 09) est imates Pakist an's adult literacy ra te at 54.9 percent (2007 figures est imated in 2008). Bo th groups of respondents exhibited a fairly uniform pattern, with the borrowing hou seholds being slight ly better off in having more lit erate a dults, as illustrated in Appendix 6.

PSLM (GoP 2009b) cap tures data across a series of indicators divided into rural and urba n categories across all four provinces, but comparison will only be made with rural Punjab, the province of this study. According to the PSLM survey, 18 percent of the total households in rural parts of Punjab have acce ss to piped wat er, 44 percent use h and pumps and 35 percent have mot orised pumps in t heir homes. These f igures were close to those obtained by the survey carried out for this study, in which 53 percent reported using hand pumps and 30 percent had mot orised pumps. Dat a published by PSL M for access t o toilet facilities revealed that 51 percent had acce ss to flushed toilet systems and 49 percent did not have any facility at all. The survey for this study found 57 per cent and 42 percent for the two classes, respectively. Dat a for drainage sy stems were ca ptured across t hree cat egories: covered, open and no facilit y, which was reported by the survey at six percent, 67 percent and 27 percent, respectively (Appendix 7).

In addition to water and sanit ation facilities, the survey for this study captured vital data relating to households' general d welling con ditions. Data collect ed for home ownership showed t hat around 94 percent of respondent s owned t he houses they were living in. Roofing structures were dominated by metal beams and bri cks at 52 p ercent, followed by wooden beams and bricks at 42 percent. Only six percent of the houses had concrete roofs. For construction of exterior walls, bricks were used in 75 p ercent of the cases, and mud for the remaining 25 per cent. Mud was more commonl y used as f looring material (68 percent) as opposed to the brick or cemen t floors found in t he remaining 32 percent of houses. Electricity for lighting was report ed at over 95 p ercent. In terms of type of energy used for cooking, the most common form was firewood (65 percent), followed by 27 percent that used animal-dung cakes (t he cheapest alt ernative); onl y eig ht percen t used me thane gas cylinders. Appendix 8 illu strates t he vario us dwe lling-related indicat ors t hat survey respondents exhibited.

Finally, the field instrument contained questions that were designed to capture elements of borrowers' behaviour, views and attitudes towards credit. As shown in the table in Appendix 9, in terms of purpose of obtaining credit, 43 percent stated that it was for establishing a new business, while 57 per cent reported its use f or expanding b usinesses. When aske d about the usef ulness of t he loan, around 81 perce int expresse d sat isfaction, but 19 percent reported not finding it beneficial. This figure of unsatisfied borrowers matches the proportion of those who had no plans f or borrowing in f uture (17 percent); around 75 percent were willing to borrow in the next cycle and around eight percent were still undecided at the time of interview. As expected, delinquency was almost absent and the repayment rate was very high (approximately 99 percent), a n indication that borro wers continue to repay regularly, despite t he difficulties that they face or t heir decision not to borrow in f uture. What is noteworthy, however, is t hat 'missed' payment ts were usually paid in the following month, and hence cannot be considered 'defaults' per se.

3. Modelling methodology

We measure the impact of treatment on the outcome, which is the impact of borrowing within MFI programmes on the livelihood of the households, by estimating the difference between individuals who received the treatment and those who did not receive the treatment. We apply the standard approach of matching widely used in the literature, formalised by Rubin (1973). This is defined as:

$$\Delta_i = Y_i^1 - Y_i^0 \tag{1}$$

where Δ_i is t he treatment effect of individual *i*, in which *i*=1,2,...,*N*. Y_i^1 and Y_i^0 are t he potential outcomes for treated and non-treated individuals, respectively. Even though we use cross-sectional data (as opposed to panel data) the equation (1) is supposed to approximate the difference between the potential outcomes before and aft er receiving t he treatment for each individual under certain assumptions. It is noted that, for each individual *i* in (1), there is only one ob served outcome and t he other is count erfactual and is not observed f rom the data. This makes it impossible to directly calculate, using cross-sectional data, the difference between the outcomes before and after treatment for each individual or household.

Therefore, equation (1) is modified to estimate the average treatment effects on the treated, $\Delta_{\tau\tau}$, which can be expressed formally as:

$$\Delta_{TT} = E(\Delta \mid D = 1) = E(Y^{1} \mid D = 1) - E(Y^{0} \mid D = 1)$$
(2)

 Δ_{TT} measures the difference between the expected outcome with and without treatment for the actual participants. The term $E(Y^1 | D = 1)$ represents expected outcomes for programme participants, while $E(Y^0 | D = 1)$ is the hypothetical outcome that would have result ed if the programme participants had *not* participated. In short, equation (2) allo ws extraction of the effect of t he treatment programme on the treated from the total effects estimated. Finally, equation (2) is used in the present study as an est imator to answer this count effectual question: 'What would be t he state of those individuals who participated in microf inance programmes if they had not actually borrowed?'

3.1 Selection bias issue

The equation (2) may be subject to selection biases, as $E(Y^0 | D = 1)$ is an unobserved counterfactual out come of t reated individu als. If t he approximation $E(Y^0 | D = 1) = E(Y^0 | D = 0)$ holds true, then non-participants can be convenient ly used as the comparison group. However, with non-experiment al data, this condition does not generally h old, since the component s which det ermine the part icipation decision also determine the out come variable of interest. Thus, the outcomes of the part icipants would differ even in the absence of programme participation, leading to selection bias.

When the bias is due to observables, we face a scenario known as *self-selection bias*. This refers to the case that the outcomes are not observed for all individuals, since they cannot

participate in the treatment programmes at the same time. One way to handle this bias is by implementing mat ching procedures, such as covariat e matching (as in Rubin 19 73) and propensity scores, as suggest ed by Rosenbaum and Rubin (1985) (RB, hereafter), which use non-participants' available information to estimate the impact In this paper, we use Propensity Score Matching (PSM) to handle the bias, since it solves the problem of multi-dimensionality which arises f rom the application of covariate matching with a large number of covariates.

Observables are not , however, t he only source of bias, as bias asso ciated wit h In ord er t o control f or this, t he inst rumental variables (I V) unobservables is possible. approach can be used, as in Heckman et al. (1997) and Mo ffitt (1996). One of the methodological advant ages in usin g st atistical ma tching rat her than t he IV estimat ion approach is that the former does not assume linearity and is valid even though distributions of explanatory variables of treatment and control groups overlap relatively little; and it does not require a valid set of instruments.³ However, the matching approach (e.g. PSM) does help to eliminat e much of an y bias asso ciated with un observables. Indeed, replicat ion studies comparing non-experimental evaluations, such as PSM, with experiments for the same programs do not appear t o have found such an example in pract ice. For example, Heckman et al. (1997) in an evalua tion of job training prog rammes have shown that the matching method applied to the control groups in the same labour markets using the same questionnaire would eliminate much of the selection bias associated with unobservables, although the remaining bias is still non-negligible. Furthermore, Chemin (2008) applied PSM to the cross-sectional household data set on Bangladesh in 199 1/92 and e valuated the impact of participation in microfinance programmes on a number of outcome indicators. The study found that microfinance had a positive impact on participants' expenditure, supply of labour and male/female school enrolment. The results are consistent with an earlier study by Pitt and Kh andker (1998), who applied the IV technique to the same data. In our data, the members of the control group were selected to be geographically close to the members of the treatment group, and the same quest ionnaire was used f or both group s, so it is conjectured that selection bias on unobservables has been minimised.

Thus, in the context of this study, the bias is defined as the difference between the outcomes of programme participation and non-participation. Formally:

$$bias = E(Y^{1} | D = 1) - E(Y^{0} | D = 0)$$
(3)

As the effect of interest of those treated participants is captured by (3), we also need to remove the effect of non-treated participants, which is defined as:

$$E(Y^{0} | D = 0) - E(Y^{0} | D = 1)$$
(4)

Equation (5) defines the sub-set of all individu als who are non-participants and h ave not been t reated. There fore t he bias is t he difference between t he effect on the t reated

³ Metho dological i ssues and p rograms for propensity sco re mat ching e stimation a re discu ssed i n detail in a number of studies, such as Becker and Ichino (2002), Dehejia (2005), Dehejia and Wahba (2002) and Smith and Todd (2005).

participants and t he difference bet ween effect s on no n-treated p articipants and non-participants. Formally:

$$\Delta_{TT} - \left[E(Y^{0} | D = 0) - E(Y^{0} | D = 1) \right] =$$

$$E(Y^{1} | D = 1) - E(Y^{0} | D = 1) - E(Y^{0} | D = 0) + E(Y^{0} | D = 1)$$

$$\Delta_{TT} - \left[E(Y^{0} | D = 0) - E(Y^{0} | D = 1) \right] = E(Y^{1} | D = 1) - E(Y^{0} | D = 0)$$
(6)

In the ideal case, the bias is zero, which implies:

$$E(Y^{1} | D = 1) - E(Y^{0} | D = 0) = 0 \iff E(Y^{1} | D = 1) = E(Y^{0} | D = 0)$$
(7)

Therefore, Δ_{TT} is ident if ident only when equat ion (7) holds, thus solving t he issue of self-selection.

3.2 PSM Estimator and estimation methodology

Equation (2) is est imated from the PSM es timator. RB int roduce what is known as a balancing score to avoid the problem of high dimensionality. The balancing score suggested by RB is defined as a propensity score, which is a function that estimates the probability of participating in the programme given the observed covariates (e.g. observed characteristics for each individual). Formally, the propensity score is defined as:

$$P(D=1 \mid X) = P(X) \tag{8}$$

This latter is estimated using one of the models available in the literature, such as the logit or probit model. These models predict the likelihood that individuals would join the microfinance programmes conditional on their personal characteristics. Following much of the literature, equation (8) is specified as a probit model and expressed as follows:

$$P(D=1 \mid X) = P(y^* > 0 \mid X) = P(u > -X\beta \mid X) = 1 - G(-X\beta) = G(X\beta)$$
(9)

where $0 < G(X\beta) < 1$, for all values of covariates X, $X\beta = \sum_{j=1}^{k} \beta_j X_j$ and G is a standard normal cumulative function. The model in (9) is non-linea r and t herefore the estimator implemented is a maximum likelihood estimator.

Equation (9) sat isfies the unconf oundness a ssumption, which implies in t his case t hat potential outcomes are independent of treatment, given t he set of covariates X such t hat: $Y^0, Y^1 \perp D \mid P(X)$, as well as t he overlap condition. This latter ensures all in dividuals with the same characteristics in the sample have a posit ive probability of being participants and non-participants (i.e. $0 < P(D = 1 \mid X) < 1$). Therefore, the PSM estimator of Δ_{TT} is selection-bias free. Formally, the PSM estimator defined is as:

$$\Delta_{TT}^{PSM} = E_{P(X)|D=1} \Big[E \Big(Y^1 \mid D = 1, P(X) \Big) - E \Big(Y^0 \mid D = 1, P(0) \Big) \Big]$$
(10)

A number of matching algorit hms have been suggest ed in t he literature to contrast the outcome of treated individuals with the outcome of individuals in the comparison group (i.e. borrowers and non-borrowers). We report the results of two matching algorithms, namely, *stratification* and *kernel* matching,⁴ which are widely used in t he literature. Using t wo matching algorithms av oids any s hortcomings t hat may result f rom relying on a single method, and it also helps to check the robustness of the estimated impact.

3.3 PSM estimates: general discussion

Appendix 10 report s the estimation output of the propensity score using t he probit model reported in the f irst panel, along with its estimated marginal effects reported in the second panel. The dependent variable is whether the household part icipated in t he microfinance programme. We assu me t hat ho usehold co mposition and charact eristics, condit ions of housing, infrastructure, and participation in t he labour market would affect the decision to participate, and we use t he redu ced f orm o f equa tion for t he prog ramme participat ion equation. T he explanat ory variables include age o f h ousehold adult s, occupat ion o f household head and adult s, child d ependency ratio, access t o electricity, home ownership status (own ed or rent ed), consumption of luxury food, such as beef , percentage of literate adults, and availability and type of toilet.

Among the explanatory variables, type of occupat ion of hou sehold head, home own ership, consumption of luxury food (bee f), and consump tion of s taple food had a nega tive and statistically significant eff ect on the likelihoo d of borro wing money, or o f joining t he programme. This implies that better living conditions as well as higher consumpt ion of beef and staple food lowered the probability of individuals join ing the programme. On the other hand, indicators such as child dep endency ratio, instances of child la bour and availability and type o f t oilet have a posit ive and st atistically significant effect on the probability of borrowing or joining the programme ; these indicators reflect t he fact that ho usehold members are in depriva tion, encouraging one of the members to borrow t o set up small family-run businesses.

Distribution of the est imated propen sity score of all t he ho useholds re sulted in some 11 observations being dropped from the matching procedure, since they lay outside the overlap region. This is shown in Appendix 11 where the propensity score dist ributions for bot h groups are displayed. Six blo cks are estimated to be within the common support region in which the balancing property is confirmed for each block and all individuals within the range [0.138, 0.982] are kept in the model. Thus, 462 borrowers are to be mat ched to 659 non - borrowers. The int ervals identified are of [0.131, 0.2], [0.2, 0.3], [0.3, 0.4], [0.4, 0.6], [0.6, 0.8], and [0. 8, 0.982] with 42, 195, 303, 512, 61 and eight overlaps in e ach block,

⁴ Stratification matching i s ba sed on splitting the p redicted prop ensity score within the common support region into intervals in a way that in each interval there are treated and controls, while Kernel matching is a non-parametric algorithm that uses weighted averages of almost all the in dividuals in the control group to construct the counterfactual outcome. See Becker and Ichino (2002) or Caliendo and Kopeinig (2008) for more details.

respectively. This gives the fourth block the largest overlap, while the last interval has the least number of individuals with common characteristics. In all blocks, the balancing property is tested and there is no significant difference between the means of treated group and control group as reported. With the balancing property satisfied and six blocks estimated, the PSM estimator satisfies the unconfoundedness and overlap conditions, and is thus bias-free.

The matching of covariates is well balanced, using the propensity score estimated within the common support region. A test of the equality (*t* statistic) of the two samples before and after matching is run f or each covariate, in which the null hy pothesis states that the means of a covariate in the comparison and treated groups are equal. If we accept the null hypothesis then the two groups are well balan ced. It has been conf irmed that all covariat es are well balanced after matching⁵ and thus matching quality for each covariate individually is not an issue.

4. Survey findings: economic and social impact of microfinance

The sections above d iscussed the methods and various procedures adopted to control the sample of any selection biases. Once tests showed that both groups (control and treatment) were at par, the average t reatment-on-treated eff ect (ATT) and t he t-statistics f or each indicator across the four dimensions of wellbeing were calculated, as shown in Ta ble 2. As discussed in det ail below, across each dimension, st atistically significant values provide strong evidence t hat disparities in bot h groups did not occur merel y by chance, but are attributable to programme participation.

⁵ Details will be furnished on request.

Table 2: Average Treatment-on-Treated effect (ATT) and *t*-statistics across various dimensions and associated indicators

Variables	KERNEL S		STRATIFICATION		Variables	KERNEL		STRATIFICATION	
variables	ATT	t-stat	ATT t-stat		ATT	t-stat	ATT	t-stat	
LIVESTOCK			1		HUMAN DEVELOPMENT INDICATORS (continued)				
Poultry 168.89		1.5	171.42	1.46	Clothing expenditure: percentage of income	-0.15	-0.66	-0.16	-0.64
Cows 4,292.73		0.89	4,096.13	0.88	Clothing expenditure: percentage of expenditure	0.48	1.64*	0.4	1.27
Total livestock value	5,241.99	1.06	4,958.42	1.07	Monthly expenditure on healthcare	148.1	3.29***	148.28	3.84***
TRANSPORT-RELATED ASSETS			Poverty ranks	0.1	2.09**	0.09	1.80*		
Motorcycle -591.33		-0.66	-896.35	-0.99	Poverty score	0.07	1.1	0.06	1.01
Bicycle 142.55		1.62	136.44	1.51	Children currently at school	0.03	0.35	0.02	0.16
Carts -231.3		-0.19	-110.98	-0.09	Monthly children's schooling expenditure	53.33	0.39	17.46	0.11
Total transport assets value	-680.08	-0.46	-870.89	-0.7	Total children in household	0.07	0.58	0.08	0.69
SAVINGS	•				Total family size	-0.02 -0.15		-0.02	-0.14
ROSCA (participation in schemes)	0.08	3.99***	0.08	4.17***	Monthly household expenditure 229		0.89	211.01	0.89
Total ROSCA encashment amount	1,722.99	1.2	1,544.77	0.94	Monthly household income	1,301.16	2.76***	1,221.75	2.60***
APPLIANCES AND ELECTRONICS			1		FOOD CONSUMPTION AND PURCHASE-RELATED INDICATORS				
Mobile phones	-104.63	-0.84	-116.35	-0.93	Consumption of luxury food: chicken 0.06 1.93*			0.05	1.62
Radio -87.57		-1.62	-83.79	-1.70*	Consumption of luxury food: mutton -0.02		-0.6	-0.02	-0.77
Sewing machine	33.01	0.32	14.66	0.15	Purchase of staple food: wheat 0.34 1.86*		1.86*	0.29	1.54
TV 364.03		1.97**	344.52	1.62	DWELLING-RELATED INDICATORS				
VCR -15.29		-0.2	-14.96	-0.21	Type of cooking fuel used -0.07 -0.98		-0.98	-0.07	-0.97
Washing machine	-65.38	-0.48	-84.09	-0.55	Material used for constructing floors 0.06 1.3		1.3	0.06	1.04
Total appliances and electronics	124.76	0.18	80.7	0.11	Overall condition of house	0.05	1.3	0.05	1.23
Value of assets per person	601.43	0.64	558.92	0.56	Material used for constructing roof	0.18	2.71***	0.17	2.53**
Total value of household assets	4,686.67	0.85	4,168.23	0.76	Material used for constructing walls	0.15	2.84***	0.15	3.06***
HUMAN DEVELOPMENT INDICATORS			Source of water supply in house	0.26	3.26***	0.23	2.64***		
Per capita expenditure on clothing and footwear	112.37	2.43**	103.35	2.08**	* Method used for waste water disposal -0.02 -0.67 -0		-0.03	-0.99	
Clothing and footwear expenses per annum	632.08	2.35**	569.86	1.90*					1

Source: Survey data. 1% *t* critical value is 2.576 (***significant at 1%). 5% *t* critical value is 1.96 (** significant at 5%). 10% *t* critical value is 1.645 (*significant at 10%).

4.1 Asset accumulation and household wellbeing

Out of the four dimensions across which various indicators were capt ured by the survey, assets tend to be more stable over time and hence are a better indicator of economic wellbeing than income or expenditure. Moreover, assets are normally calculated to represent an annual est imate and represent the enduring results of income f lows and expenditures. Another important role that household assets play during 'lean' period s is helping to cope with adverse condit ions and in periods of low and unst able income; as their disposal can 'smooth' consumption and expendit ure during crises. Household assets in the survey were captured across two dimensions: physical assets (tangible) and human capital (intangible). Tangible ho usehold assets were f urther classified into livest ock, transport-related assets, savings (financial capital), and appliances and electronics.

Livestock const itute an import ant category of assets for the rural poor, as they can be classified as 'income-generating' and provide a means of livelihood. A substantial portion of borrowing was done to purchase cows and goats, and some household's relied exclusively on livestock as a source of income , al though they were found to provide supplement ary income in most cases. Survey findings show that borrowers seem to fare better in terms of livestock-related assets, albeit not to a significant level. Differences in poultry, being of small monetary value, show borrowers to be marginally at an advantage (on the average between both methods) by around Rs.170; they were statistically non-significant with t statistics 1.50. ATT for cows was positive and large, but not statistically significant and does not lead to any firm conclusion.

In the case of transport-related assets, non-borrowers seemed t o fare better, although the differences were not statistically significant. Bicycles were t he only as set where borrowers seemed t o be bet ter o ff, b y small amount s, as compared t o non-borrowers, by values ranging from Rs.136 to Rs.142 across the two methods used for comparison, with t statistics ranging from 1.51 to 1.62.

Savings constitute an important component of financial capital. Robinson (2001:21) argues that:

deposit services are more valuable than credit for poorer h ouseholds. With saving s, not only can households build up asset s to use as collateral, but they can also better smooth seasonal consumption needs, finance major expenditures such as school fees, self-insure against major shocks, and self-finance investments.

Owing to the variation in policies and the erratic and inconsistent saving behaviour of client households, the most suitable and relevant proxy for establishing saving behaviour of respondents was considering part icipation in ROSCA (Rotating Savings and Credit Association) schemes, which are a form of informal saving model found in many parts of the world, known by different names. Survey findings show that there is a marked difference in saving behaviour across both groups. As shown in Table 2, borrowers show a much higher probability and inciden ce of part icipation in ROSCA schemes t han did non-borrowers. Moreover, there was a n average difference (ranging f rom Rs. 1,723 to Rs. 1,545, across kernel and stratification methods) in the encashment amount of the scheme, with borrowers

saving grea ter a mounts and, as would be expect ed, con tributing mo re (around Rs. 105 monthly) towards instalments. A possible explanation is that once rural households start to participate in microcredit programmes, they develop a sense of financial access and realise the importance of participating in saving sche mes. In the absence of formal options, they resort to semi-formal models (such as ROSCA, in this case) and commit a certain amount to be contributed.

As opposed to livestock, the impact of borrowing on applia nces and electronics was not so pronounced. There was a very s mall, almost negligible diff erence across household electronics such as fridges, VCRs and sewing machines, whereas non-borrowers seemed to fare slightly better in terms of owning radios. Borrowers, however, seemed to be better off in owning televisions (with average difference in values ranging from Rs.344 to Rs.364 across both methods) as compared to non-borrowers. Borrowers were also f ound to be better off if comparisons were made of the overall value of appliances and elect ronics, alt hough the difference was not statistically significant. The overall value of total or per capit a household tangible assets owned by borrowers was found to be greater than that of non- borrowers, but it is not statistically significant.

4.2 Human resources

Our surve y questionnaire also capt ures various demographic characteristics of household members, household income and amoun ts pent on cl othing and f ootwear, children's schooling, and healt hcare. Clot hing and f ootwear expendit ure shows t hat borrower households spend more t han non-borrowers, and t he difference ranges f rom Rs. 569 to Rs.632 which is statistically significant at the five percent level. Calculations also reveal that borrowing h ouseholds' spending on healt hcare was on a verage Rs. 148 more than non-borrowers' and the difference is statistically significant at the one percent level. In terms of indicators on literacy, borrowing households were found to be slightly better in terms of adult literacy, while school at tendance was found to be almost the same for both groups. There was, howe ver, a small and non-sig nificant diff erence in t he amount of average mon thly schooling e xpenditure with borrower households spend ing more. There are minor , almost negligible, diff erences when households are compared for total adults, children and total family size.

4.3 Household income and expenditure

Table 2 port rays the differences between both groups of respondents in terms of monthly household income and expenditure. While the difference in expenditure is inconse quential (varying from Rs.211 to Rs.230 across matching methods), the difference in income is both substantial (given t hat the sample's median in come is Rs. 7,500), as well a s st atistically significant at the one percent level. Depending on t he matching method used, mon thly income of borrowers is greater by Rs.1,221 (stratification) and Rs. 1,301 (kernel m ethod). This disparity can be att ributed to a number of factors. One possible explanation is t hat borrowers supplement their income b y ob taining microcredit and invest ing the a mount in livestock or other small income-generating as sets, such as a sewin g machine, bicycle or cart. On the other hand, if they have access to savings, borrowers can combine credit from

the MFI and invest in a larger asset, which acts as the primary source of income. Examples from the survey include setting up a roadside hotel, a barber's shop, a bicy cle repair shop, buying a donkey-cart, purchasing a cow or selling an existing one and 'upgrading' to a better breed.

4.4 Food security and consumption behaviour

The present study focuses on dietary diversity, food quality, frequency of purchase and stock of storable staple foods as proxy indicators for food security. As shown in the calculations, borrowers were seen to fare better in terms of consuming 'luxury food' (chicken) more often than non-borrowers. The indicator was captured by enquiring how many days the household consumes chicken or mutt on (both identified as luxury foods within the local cont ext). For ease of recall and to ensure accur acy, the period was kept to one week. The frequency of chicken con sumption was f ound to be sig nificant (at the 10 percent level), while mutt on favoured non-borrowers by a negligibly small amount. Since borrowing households consume more luxury foods, consumption of staple food (wheat, in the case of this survey) was found to occur in greater frequency amongst non-borrowing households, as would be expected.

Other indicators in this dimension were the frequency of purchase and the stocks of storable staple food held on the premises. These indicators are very sensitive and capt ure relative household wellbeing by estimating the number of weeks of wheat that the household has in store, the p roxy for which was the frequency of its purchase. Poorer household s were observed to purchase more frequently, possibly due to liquidity constraints, with the poorest having to purchase on a daily basis. The frequency was capt ured across an ordered variable, ranging from a daily basis to weekly, fortnightly, monthly, biannually and annually. Table 2 shows that borrowers seem to be better off in terms of holding stocks of wheat, as the *purchase of wheat* indicator was found to be statistically significant (at the 10 percent level).

4.5 Dwelling-related indicators

The dimension t hat me asured hou sing conditions was captured across various indicat ors, such as t he type o f cooking f uel used, en ergy used for ligh ting, ma terial u sed f or constructing f loors, roo fs, walls, source of water supply, and the method used for waste water disposal. Finally, the overall condition of the house was ranked during interviews by observing its condition. The result s show t hat borrowers seem to live in bett er conditions than non-borrowers a cross all indicat ors, except for the type of cooking fuel used and the method of disposing of wast e water, where non-borrower s sho w ver y slight, ne gligible instances of being a t an advant age. The most pronou nced and st atistically s ignificant differences were found in 'the type and mat erial used for constructing roofs, internal and external walls' and 'the source of water supply in the house'. All of t hese reflect be tter dwelling conditions enjoyed by borrowers.

5. Concluding remarks

Drawing up on a primary provincia I-level cro ss-sectional household survey cond ucted in Pakistan, the present st udy anal yses t he ext ent and direct ion of pro gramme i mpact on borrowers, assessed through a range of dimensions that captured and reflected relative wellbeing of a typical rural household in Pakist an. Household charact eristics were captured across four dimensions, further segregated into various ind icators, the data on which wa s gathered b y administering a semi-st ructured quest ionnaire in t he field. The research wa s based on the quasi-experimental design that compared differences bet ween borrowers and non-borrowers. I n ord er t o cont rol f or an y selection bias t hat may have arisen during sampling of households, the propensity score matching model was applied, through which the average treatment-on-treated effect was finally computed.

As discussed in the previous sections, borrowers were seen to fare better in most of the indicators across vario us dimensions of relative household wellbeing. The ext ent of the difference across bot h groups was subst antial as well as st atistically significant in some indicators, while it was found to be weak and negligible in others. For example, borrowers performed better in terms of livestock, participation in savings schemes, and overall value of household assets. Borrowers' household income and expenditure was also seen to be better and in terms of food consumption they had a slight edge over non-borrowers, as they were found to consume more 'luxury' foods and also had larger stocks of storable staple foods. In the case of dwelling-related indicators, borrowers had a b etter quality of floors, roofs, walls, and wat er supply in t heir house s, alt hough non-borrower s seemed t o use bet ter quality cooking fuel and had improved wast e water disposal systems. The most prominent and statistically significant differences across bot h groups favoured b orrowers, and were observed in savings, t elevisions, expendit ure on healt hcare, monthly household income, expenditure on clot hing and f ootwear, and certain dwe lling characteristics, such a s water supply and quality of roofing and walls. Overall, borrowers were seen to be better in around 70 percent of the indicat ors across which comparisons were made in t he final model. Borrowing households, in comparison wit h non-borrowers, were t herefore able t o increase household income by investing more in prod uctive assets, such as livestock or sewing machines; this income was either saved for future investment or was consumed in the form of 'luxury' foods or for stocking staple food items, or was in curred on healthcare. Given the persistence of poverty and vulnerability in rural Pakistan, the results show that microfinance can be used as an effective measure in alleviating poverty in the country.

As the nature of poverty is multi-dimensional, people's needs are unique and hence have to be addressed by o ffering them dist inctive, cu stomised solutions. MFIs in Pakist an lack innovation and have a limited number of programmes to offer. The 'one size fits all' approach was observed across a Imost all le nders who formed part of the survey, as most of them offered basic credit and saving facilities, with rigid rules regarding interests rates, loan sizes, or borrower selection criteria. Most of the successful MFIs in the world have been observed to have an assortment of products and services that are tailor-made to suit specific groups of vulnerable clients. BRAC's progra mmes committed to targeting the ultra poor (TUP and IGVGD) an d Gra meen Bank's beggar loans are such examples. These programmes combine livelihood prot ection (food aid, emplo yment) with liveliho od promotion (financial services with skills training) and are geared towards assisting the poorest to gradually move out of poverty. Pakistan would need to implement programmes such as these to address the multi-dimensional poverty and bring about real change to livelihoods.^{6 7}

Despite the limitations in the methodology of PSM applied to cross-sectional data, such as the possible bias arising f rom u nobservable f actors, the st udy has conf irmed t hat microfinance programmes had a po sitive impact on the welfare of participating households; that is, the poverty-reducing effects were observed and statistically significant on a number of indicat ors, including expenditure on healt hcare or clot hing, mon thly household income, and certain dwelling ch aracteristics, such as wat er supply and quality of roofing and walls. Much more sustained efforts, how ever, such as t ailoring the microfinance programmes to meet borro wers' dema nds, would make t he positive impact more pro nounced, given t he limited access to financial services in Pakistan.

⁶ During focus groups and individual interviews, many borrowers complained of the size of the loan, which was too small to st art any busi ness and re quired very frequent repayments. If lenders are sensitive to su ch basic borrower demands, the impact will be more pronounced without affecting institutional sustainability.

⁷ Limited access to financi al services in the devel oping world is o ne of the main obsta cles to both income generation and social protection. Demirgüç-Kunt et al. (2 008) use a composite measure of estimating financial inclusion and reveal that only 12 percent of people in Pakistan have access to an account with a financial intermediary. This is seen to be especially low if com pared to 48 percent in India, 59 percent in Sri Lanka, and 32 percent in Bangladesh (Haq 2008).

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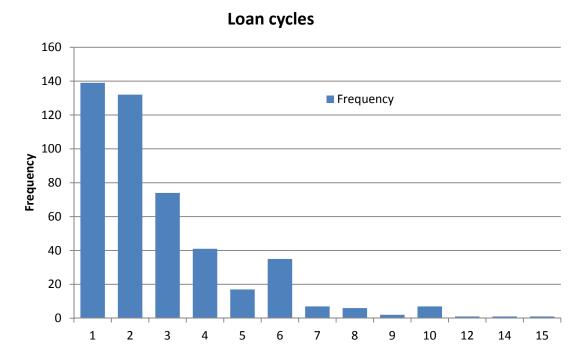
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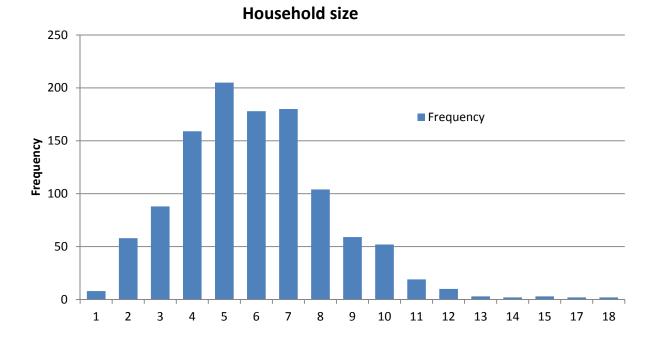
Appendix 1: Distribution of institutional participation among survey
participants

Microfinance institution	1	2	3	%	Grand
					total
National Rural Support Programme (NRSP)	153	4	1	31.66	158
Kashf Foundation	138	2	0	28.06	140
Punjab Rural Support Programme (PRSP)	67	2	2	14.23	71
Khushhali Bank	39	0	0	7.82	39
Pak Oman Bank	25	0	0	5.01	25
CSC	22	8	3	6.61	33
1 st Microfinance Bank	13	2	1	3.21	16
Asasah	6	8	3	3.41	17
Total	463	26	10	100	499



Appendix 2: Distribution of survey respondents showing loan cycles completed

Number of loan cycles completed



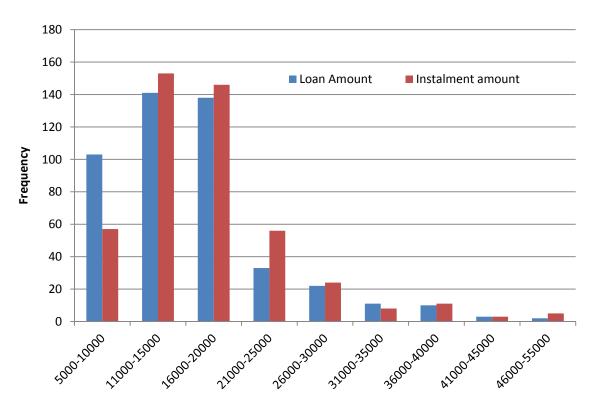
Appendix 3: Distribution of household size among survey participants

Number of household members

Loan Amount (Pakistani Rupees)	Frequency	Percentage	Instalment amount (Pakistani Rupees)	Frequency	Percentage
5000-10000	103	22.25	0-1000	57	12.31
11000-15000	141	30.45	1001-1500	153	33.05
16000-20000	138	29.81	1501-2000	146	31.53
21000-25000	33	7.13	2001-2500	56	12.10
26000-30000	22	4.75	2501-3000	24	5.18
31000-35000	11	2.38	3001-3500	8	1.73
36000-40000	10	2.16	3501-4000	11	2.38
41000-45000	3	0.65	4001-4500	3	0.65
46000-55000	2	0.43	4501-5500	5	1.08
Total	463	100	Total	463	100

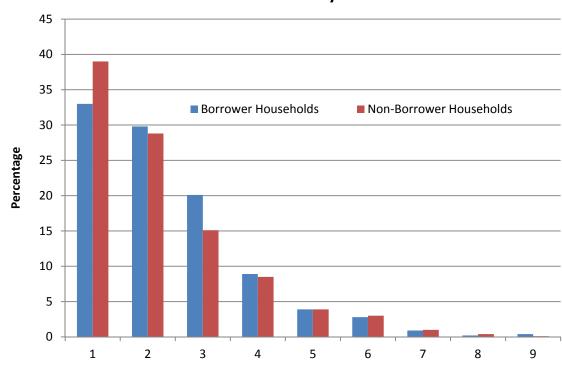
Appendix 4: Loan sizes and instalment amounts of borrowers interviewed





Loan and instalment amounts

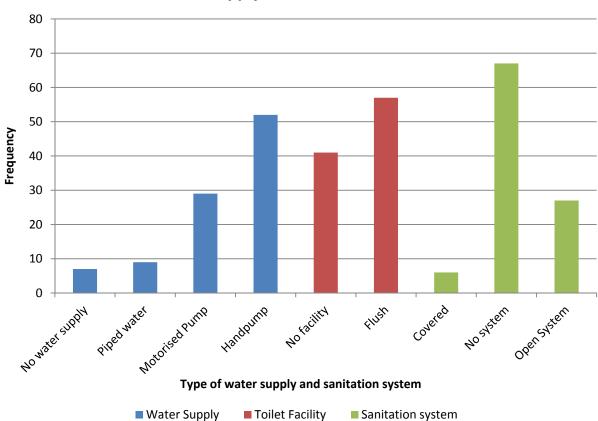
Amount (Pakistani rupees)



Adult literacy

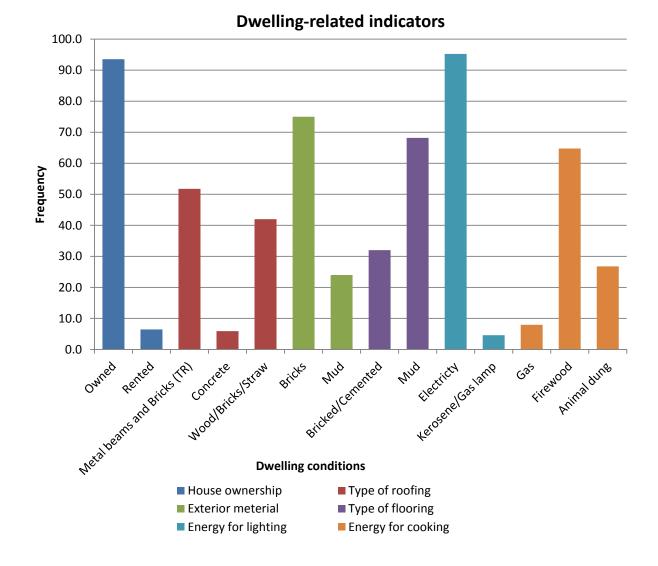
Appendix 6: Comparison of borrower and non-borrower households for adult literacy

Number of literate adults in household



Water supply and sanitation facilities

Appendix 7: Types of water supply and sanitation facilities available to survey respondents



Appendix 8: General dwelling conditions of surveyed respondents

Appendix 9: Basic indicators showing loan use and satisfaction among survey participants

Indicators	Frequency	Percentage
Purpose of obtaining credit	1	
New business	202	43.63
Expansion	261	56.37
Was the loan beneficial?	1	
Yes	375	80.99
No	88	19.01
Plans for future borrowing	1	
Yes	346	74.73
No	80	17.28
Not sure/will think about it	37	7.99
Missed payments		
No	458	98.92
Yes	5	1.08

Appendix 10: LPM and Probit estimated score (Dependent variable: whether a household participated in the microfinance programme)

Variables	E	Probit stimates	Probit Marginal Effects		
	β	p – value	β	p – value	
Intercept 1.662		0.011	-	-	
Value of agricultural land	0.008	0.936 0.003	0.	936	
Average age of household adults	0.006	0.252 0.002	0.	252	
Type of occupation of household head	-0.088	0.017 -0.034	0.	017	
Child dependency ratio	0.098	0.030 0.038	0.	030	
Child labour	0.206	0.021 0.080	0.	021	
Elect electricity supply in house	-0.227	0.216 -0.088	0	216	
Value of goats/sheep	0.000	0.009 0.000	0.	009	
Home ownership status (owned or rented)	-0.465	0.008 -0.180	0	008	
Consumption of luxury food: beef	-0.233	0.031 -0.090	0.	031	
Occupation of adults	-0.050	0.129 -0.019	0	129	
Percentage of literate adults	0.002	0.093 0.001	0.	093	
Number of rooms in house	-0.030	0.400 -0.012	0	400	
Consumption of staple food	-0.196	0.010 -0.076	0	010	
Availability and type of toilet	0.174	0.028 0.068	0.	028	
Stock of wheat held	-0.003	0.155 -0.001	0.	155	
N <i>†</i>		1127		1127	

 β : refers to estimated coefficients.

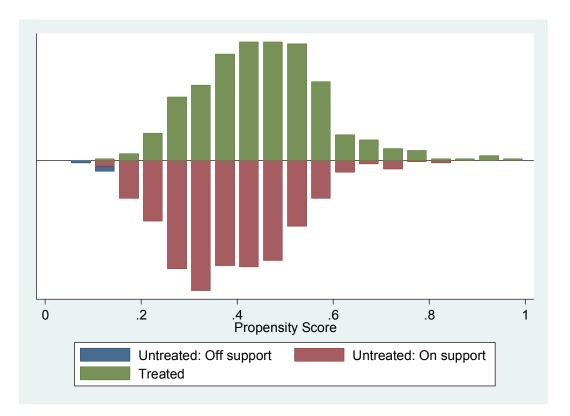
: The test statistics for the estimated probit model is based on the standard normal distribution, unlike the linear probability model that is based on the t distribution.

 τ : *N*: is the number of observations.

LR is the log likelihood ratio estimated for the probit model. Both statistics are to test the null hypothesis that states the model is jointly is not significant. If the hypothesis is accepted then the model is overall not significant, which implies the set of covariates need to be changed. Values between parentheses are p values.

p. R^2 : pseudo R^2 is the goodness of fit measure estimated for the probit model.







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